Generalized pairwise comparisons for right-censored time to event outcomes

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Clinical trials in oncology

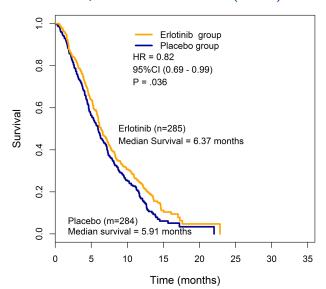
Efficacy/safety can reveal opposite effects of the treatment:

- longer survival
- serious but non-lethal adverse effects

However, efficacy and safety outcomes are usually analyzed and reported independently

- efficacy: using log-rank test
- safety: using Fisher's exact test

Example - Moore et al. (2007)



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Worst grade related adverse event	Erlotinib group (n=282)	Placebo group (n=280)
Grade 1	48 (17%)	69 (24.6%)
Grade 2	118 (41.8%)	89 (31.5%)
Grade 3	72 (25.5%)	47 (16.8%)
Grade 4	11 (3.9%)	6 (2.1%)
Grade 5	4 (1.4%)	3 (1.1%)

Clinical trials - handling multiple endpoints

I don't think there is a good objective approach.

What about a good subjective approach?

Patient preference

- 1. increase survival by at least 2 months
- 2. otherwise, least serious adverse events

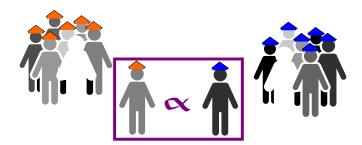
Generalized pairwise comparisons (GPC)

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GPC - 1 endpoint

Consider N patients divided into two groups:

- experimental group: m patients with response $(x_i)_{i \in \{1,\dots,m\}}$
- control group: n patients with response $(y_i)_{i \in \{1,...,n\}}$

Denote by $\tau \in \mathbb{R}^{+*}$ the smallest difference in response that is clinically relevant.

Our parameter of interest is the net benefit:

$$\Delta = \mathbb{P}\left[X \geq Y + \tau\right] - \mathbb{P}\left[Y \geq X + \tau\right]$$

Point estimation in GPC

Defining $s_{ij} = \mathbb{1}_{x_i \geq y_i + \tau} - \mathbb{1}_{y_i \geq x_i + \tau}$ the score of the pair i, j:

GPC 0

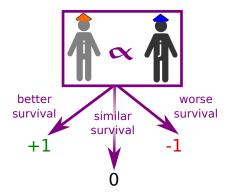
$$\hat{\Delta} = \frac{1}{nm} \sum_{i=1}^{n} \sum_{j=1}^{m} s_{ij}$$



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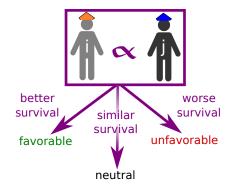
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Point estimation in GPC

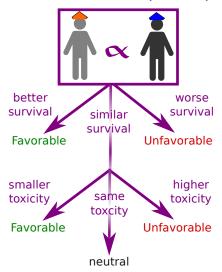
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GPC

Generalization to multiple endpoints



i.e. $s_{ii} = \phi(\mathbf{x}_i, \mathbf{y}_i, \boldsymbol{\tau}) - \phi(\mathbf{x}_i, \mathbf{y}_i, \boldsymbol{\tau})$ where ϕ is a scoring rule.

GPC in presence of censoring

We only observe:

- right-censored event times $(\tilde{x}_i)_{i\in\{1,\dots,m\}}$ and $(\tilde{y}_j)_{j\in\{1,\dots,n\}}$
- event type indicators $(\delta_{{\sf x},i})_{i\in\{1,\dots,m\}}$ and $(\delta_{{\sf y},j})_{j\in\{1,\dots,n\}}$.

How can we compute $s_{ij} = \mathbb{1}_{x_i \geq y_j + \tau} - \mathbb{1}_{y_j \geq x_i + \tau}$?

Set s_{ij} to 0 when the pair cannot be decidedly classified.

Gehan scoring rule

GPC in presence of censoring

Set s_{ii} to 0 when the pair cannot be decidedly classified.

If $\delta_{x,i} = 0$ and $\delta_{y,i} = 1$:

Example 1: $\tilde{x}_i \geq \tilde{y}_i + \tau$

• $s_{ij}=1$

Example 2: $\tilde{x}_i < \tilde{y}_i + \tau$

• $s_{ii} = 0$: uninformative pair

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riangle uninformative pairs bias the estimation of Δ toward the null (Buyse, 2008)

Peron scoring rule - Péron et al. (2018)

GPC in presence of censoring

Compute

- $p_{ii}^f = \mathbb{P}\left[X \geq Y + \tau \middle| X \geq \tilde{x}_i, X \geq \tilde{y}_i, \delta_{x,i}, \delta_{y,j}\right]$
- $p_{ii}^{uf} = \mathbb{P}\left[Y \geq X + \tau | X \geq \tilde{x}_i, X \geq \tilde{y}_i, \delta_{x,i}, \delta_{y,j}\right]$
- $s_{ij} = p_{ii}^f p_{ii}^{uf}$

Peron scoring rule - Péron et al. (2018)

Compute

•
$$p_{ij}^f = \mathbb{P}\left[X \geq Y + \tau \middle| X \geq \tilde{x}_i, X \geq \tilde{y}_j, \delta_{x,i}, \delta_{y,j}\right]$$

•
$$p_{ij}^{uf} = \mathbb{P}\left[Y \ge X + \tau \middle| X \ge \tilde{x}_i, X \ge \tilde{y}_j, \delta_{x,i}, \delta_{y,j}\right]$$

•
$$s_{ij} = p_{ij}^f - p_{ij}^{uf}$$

If $\delta_{x,i} = 0$ and $\delta_{y,i} = 1$

•
$$p_{ij}^f = \min\left(\frac{S_X(\tilde{y}_j + \tau)}{S_X(\tilde{x}_i)}, 1\right)$$

where S_x is the survival in the experimental group.

Peron scoring rule - Péron et al. (2018)

GPC in presence of censoring

Compute

•
$$p_{ij}^f = \mathbb{P}\left[X \geq Y + \tau \middle| X \geq \tilde{x}_i, X \geq \tilde{y}_j, \delta_{x,i}, \delta_{y,j}\right]$$

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•
$$s_{ij} = p_{ij}^f - p_{ij}^{uf}$$

If
$$\delta_{x,i} = 0$$
 and $\delta_{y,i} = 1$

•
$$p_{ij}^f = \min\left(\frac{S_X(\tilde{y}_j + \tau)}{S_X(\tilde{x}_i)}, 1\right)$$

where S_x is the survival in the experimental group.

Example 1: $\tilde{x}_i > \tilde{y}_i + \tau$

•
$$s_{ij}=1$$

Example 2:
$$\mathbb{P}\left[X \geq \tilde{y}_i + \tau | X \geq \tilde{x}_i\right] = 1$$

•
$$s_{ij} = 1$$

Implementation

Package BuyseTest (available on CRAN):

• S_X and S_Y estimated using Kaplan-Meier

Example of code:

```
ff <- group \sim tte(survival, censoring = event, threshold = 2) +
    cont(toxicity)
e.Peron <- BuyseTest(formula = ff,
                     data = dt.fol,
                     scoring.rule = "Peron")
```

Implementation

Example of output:

```
summary(e.Gehan)
```

Generalized pairwise comparisons with 2 prioritized endpoints

```
[...]
> treatment groups: gemcitabine (control) vs. folfirinox (treatment)
[...]
endpoint threshold total favorable unfavorable neutral uninf delta Delta
survival 2 100.00 44.74 20.97 12.50 21.79 0.2377 0.2377
toxicity 1e-12 34.29 14.50 8.53 11.27 0.00 0.0597 0.2975
```

```
summary(e.Peron)
```

Generalized pairwise comparisons with 2 prioritized endpoints

```
[...]
endpoint threshold total favorable unfavorable neutral uninf delta Delta
survival 2 100.00 56.88 26.49 16.61 0.02 0.3039 0.3039
toxicity 1e-12 16.63 6.53 4.41 5.69 0.00 0.0212 0.3251
```

Limitations & perspectives

The Peron scoring rule requires a consistent estimator for the survival

- e.g. at $\tilde{y}_i + \tau$: may not be available
- remaining uninformative pairs

Ideas:

- lower and upper bound for p^f and p^{uf} (implemented)
- parametric model for S_X and S_Y
- inverse probability of censoring weights

Reference I

- Buyse, M. (2008). Reformulating the hazard ratio to enhance communication with clinical investigators. *Clinical Trials*, 5(6):641.
- Moore, M. J., Goldstein, D., Hamm, J., Figer, A., Hecht, J. R., Gallinger, S., Au, H. J., Murawa, P., Walde, D., Wolff, R. A., et al. (2007). Erlotinib plus gemcitabine compared with gemcitabine alone in patients with advanced pancreatic cancer: a phase iii trial of the national cancer institute of canada clinical trials group. *Journal of clinical oncology*, 25(15):1960–1966.
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